Probabilistic Modeling and Expectation Maximization

CMSC 678 UMBC

Outline

Latent and probabilistic modeling Generative Modeling Example 1: A Model of Rolling a Die Example 2: A Model of Conditional Die Rolls

EM (Expectation Maximization) Basic idea Three coins example Why EM works

So far, we've (mostly) had *labeled* data pairs (x, y), and built classifiers p(y | x)

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What if we want to model *both* x and y together?

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A: Linear Discriminant Analysis

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What if we want to model *both* x and y together?

p(x, y)

Q: Where have we used p(x,y)?

A: Linear Discriminant Analysis

Like A3, Q1

Piazza Q68

•

Or what if we only have data but no labels?

p(x)

Generative Stories

"A useful way to develop probabilistic models is to tell a generative story. This is a *fictional* story that explains how you believe your training data came into existence." --- CIML Ch 9.5

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Generative stories are most often used with joint models p(x, y).... but despite their name, generative stories are applicable to both generative and conditional models

p(x, y) vs. p(y | x): Models of our Data

p(x, y) is the **joint** distribution

Two main options for estimating:

1. Directly

2.

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- 1. Directly
- 2. Using Bayes rule: p(x, y) = p(x | y)p(y)

Using Bayes rule *transparently* provides a **generative story** for how our data x and labels y are generated

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1. Directly: used when you *only* care about making the right prediction

Examples: perceptron, logistic regression, neural networks (we've covered)

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Examples: perceptron, logistic regression, neural networks (we've covered)

2. Estimate the joint

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Example: Rolling a Die

$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2)\cdots p(w_N) = \prod_i p(w_i)$$

Example: Rolling a Die

N different
(independent) rolls
$$p(w_1, w_2, ..., w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

 $w_1 = 1$ •
 $w_2 = 5$ ••
 $w_3 = 4$ ••

. . .

N different
(independent) rolls
$$p(w_1, w_2, ..., w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

$$w_1 = 1$$

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$$w_3 = 4$$

. . .

Generative Story for roll i = 1 to N:

N different
(independent) rolls
$$p(w_1, w_2, ..., w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

$$w_1 = 1$$

$$w_2 = 5$$

$$w_3 = 4$$

. . .

Generative Story for roll i = 1 to N: $w_i \sim Cat(\theta)$



• •



$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

maximize (log-) likelihood to learn the probability parameters

Q: Why is maximizing loglikelihood a reasonable thing to do?

$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

maximize (log-) likelihood to learn the probability parameters

Q: Why is maximizing loglikelihood a reasonable thing to do?

A: Develop a good model for what we observe

$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

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Q: Why is maximizing loglikelihood a reasonable thing to do?

Q: (for discrete observations) What loss function do we minimize to maximize log-likelihood? A: Develop a good model for what we observe

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maximize (log-) likelihood to learn the probability parameters

Q: Why is maximizing loglikelihood a reasonable thing to do?

Q: (for discrete observations) What loss function do we minimize to maximize log-likelihood? A: Develop a good model for what we observe

A: Cross-entropy

$$p(w_1, w_2, ..., w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

maximize (log-) likelihood to learn the probability parameters

If you observe these 9 rolls... ...what are "reasonable" estimates for p(w)?



p(5) = ? p(6) = ?

$$p(w_1, w_2, ..., w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

maximize (log-) likelihood to learn the probability parameters

If you observe these 9 rolls... ...what are "reasonable" estimates for p(w)?



$$p(1) = 2/9$$
 $p(2) = 1/9$
 $p(3) = 1/9$ $p(4) = 3/9$ maximum
likelihood
estimates
 $p(5) = 1/9$ $p(6) = 1/9$

N different
(independent) rolls
$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

$$w_1 = 1$$

$$w_2 = 5$$

$$w_3 = 4$$

Generative Story for roll i = 1 to N: $w_i \sim Cat(\theta)$

Maximize Log-likelihood

$$\mathcal{L}(\theta) = \sum_{i} \log p_{\theta}(w_{i})$$
$$= \sum_{i} \log \theta_{w_{i}}$$

N different (independent) rolls $p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$

Generative Story

for roll i = 1 to N: $w_i \sim Cat(\theta)$ Maximize Log-likelihood

$$\mathcal{L}(\theta) = \sum_{i} \log \theta_{w_i}$$

Q: What's an easy way to maximize this, as written *exactly* (even without calculus)?

N different (independent) rolls $p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$

Generative Story for roll i = 1 to N: $w_i \sim Cat(\theta)$ Maximize Log-likelihood

$$\mathcal{L}(\theta) = \sum_{i} \log \theta_{w_i}$$

Q: What's an easy way to maximize this, as written *exactly* (even without calculus)?

A: Just keep increasing θ_k (we know θ must be a distribution, but it's not specified)

N different
(independent) rolls
$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

Maximize Log-likelihood (with distribution constraints)

$$\mathcal{L}(\theta) = \sum_{i} \log \theta_{w_i} \text{ s.t.} \sum_{k=1}^{6} \theta_k = 1 \quad 0$$

(we can include the inequality constraints $0 \le \theta_k$, but it complicates the problem and, *right now*, is not needed)

solve using Lagrange multipliers

N different
(independent) rolls
$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

Maximize Log-likelihood (with distribution constraints)

$$\mathcal{F}(\theta) = \sum_{i} \log \theta_{w_i} - \lambda \left(\sum_{k=1}^{6} \theta_k - 1 \right)$$

(we can include the inequality constraints $0 \le \theta_k$, but it complicates the problem and, *right now*, is not needed)

$$\frac{\partial \mathcal{F}(\theta)}{\partial \theta_k} = \sum_{i:w_i=k} \frac{1}{\theta_{w_i}} - \lambda \qquad \frac{\partial \mathcal{F}(\theta)}{\partial \lambda} = -\sum_{k=1}^6 \theta_k + 1$$

N different
(independent) rolls
$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

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(we can include the inequality constraints $0 \le \theta_k$, but it complicates the problem and, *right now*, is not needed)

$$\theta_k = \frac{\sum_{i:w_i=k} 1}{\lambda}$$
 optimal λ when $\sum_{k=1}^6 \theta_k = 1$

N different
(independent) rolls
$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

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(we can include the inequality constraints $0 \le \theta_k$, but it complicates the problem and, *right now*, is not needed)

$$\theta_k = \frac{\sum_{i:w_i=k} 1}{\sum_k \sum_{i:w_i=k} 1} = \frac{N_k}{N} \quad \text{optimal } \lambda \text{ when } \sum_{k=1}^6 \theta_k = 1$$

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Example: Conditionally Rolling a Die

$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

$$\int_i^{add \ complexity \ to \ better}_{explain \ what \ we \ see}$$

$$p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1) \cdots p(z_N)p(w_N|z_N)$$

$$= \prod_i p(w_i|z_i) p(z_i)$$

Example: Conditionally Rolling a Die

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$$p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1) \cdots p(z_N)p(w_N|z_N)$$

$$= \prod_i p(w_i|z_i) \ p(z_i)$$

First flip a coin...

$$\sum_{i=1}^{\infty} z_1 = T$$
$$\sum_{i=1}^{\infty} z_2 = H$$
$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2)\cdots p(w_N) = \prod_i p(w_i)$ add **complexity** to better explain what we see $p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1)\cdots p(z_N)p(w_N|z_N)$ $= \prod_{i} p(w_i | \mathbf{z}_i) p(\mathbf{z}_i)$...then roll a different die First flip a coin... depending on the coin flip $z_1 = T \quad w_1 = 1$ $z_2 = H \quad w_2 = 5$

$$p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i)$$

$$\prod_{\substack{add \text{ complexity to better}\\explain what we see}} p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1) \cdots p(z_N)p(w_N|z_N)$$

$$= \prod_i p(w_i|z_i) p(z_i)$$

If you observe the z_i values, this is easy!

$$p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

If you observe the z_i
values, this is easy!

First: Write the Generative Story

 $\lambda = \text{distribution over coin } (z)$ $\gamma^{(H)} = \text{distribution for die when coin comes up heads}$ $\gamma^{(T)} = \text{distribution for die when coin comes up tails}$ for item i = 1 to N: $z_i \sim \text{Bernoulli}(\lambda)$ $w_i \sim \text{Cat}(\gamma^{(z_i)})$

$$p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

If you observe the z_i
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First: Write the Generative Story $\lambda = \text{distribution over coin } (z)$ $\gamma^{(H)} = \text{distribution for H die}$ $\gamma^{(T)} = \text{distribution for T die}$ for item i = 1 to N: $z_i \sim \text{Bernoulli}(\lambda)$ $w_i \sim \text{Cat}(\gamma^{(z_i)})$ Second: Generative Story \rightarrow Objective

$$\mathcal{F}(\theta) = \sum_{i}^{n} (\log \lambda_{z_i} + \log \gamma_{w_i}^{(z_i)})$$

Lagrange multiplier constraints

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

If you observe the z_i
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First: Write the Generative Story $\lambda = \text{distribution over coin } (z)$ $\gamma^{(H)} = \text{distribution for H die}$ $\gamma^{(T)} = \text{distribution for T die}$ for item i = 1 to N: $z_i \sim \text{Bernoulli}(\lambda)$ $w_i \sim \text{Cat}(\gamma^{(z_i)})$ Second: Generative Story \rightarrow Objective

$$\mathcal{F}(\theta) = \sum_{i}^{n} (\log \lambda_{z_i} + \log \gamma_{w_i}^{(z_i)})$$
$$-\eta \left(\sum_{k=1}^{2} \lambda_k - 1\right) - \sum_{k=1}^{2} \delta_k \left(\sum_{j=1}^{6} \gamma_j^{(k)} - 1\right)$$

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

If you observe the z_i But if you don't observe the values, this is easy! z_i values, this is not easy!

First: Write the Generative Story $\lambda = \text{distribution over coin } (z)$ $\gamma^{(H)} = \text{distribution for H die}$ $\gamma^{(T)} = \text{distribution for T die}$ for item i = 1 to N: $z_i \sim \text{Bernoulli}(\lambda)$ $w_i \sim \text{Cat}(\gamma^{(z_i)})$ Second: Generative Story \rightarrow Objective

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$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

goal: maximize (log-)likelihood we don't actually observe these z values we just see the items w

if we *did* observe *z*, estimating the probability parameters would be easy... but we don't! :(

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

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if we *did* observe *z*, estimating the probability parameters would be easy... but we don't! :(if we *knew* the probability parameters then we could estimate *z* and evaluate likelihood... but we don't! :(

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

we don't actually observe these *z* values

goal: maximize *marginalized* (log-)likelihood

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i \left[p(w_i | z_i) p(z_i) \right]$$

we don't actually observe these *z* values

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W

$$p(z_1, w_1, z_2, w_2, ..., z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

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goal: maximize *marginalized* (log-)likelihood



$$p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = \prod_i p(w_i | z_i) p(z_i)$$

we don't actually observe these *z* values

goal: maximize *marginalized* (log-)likelihood



$$p(w_1, w_2, \dots, w_N) = \left(\sum_{z_1} p(z_1, w_1)\right) \left(\sum_{z_2} p(z_2, w_2)\right) \cdots \left(\sum_{z_N} p(z_N, wN)\right)$$

 $p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1) \cdots p(z_N)p(w_N|z_N)$

goal: maximize *marginalized* (log-)likelihood



$$p(w_1, w_2, \dots, w_N) = \left(\sum_{z_1} p(z_1, w_1)\right) \left(\sum_{z_2} p(z_2, w_2)\right) \cdots \left(\sum_{z_N} p(z_N, wN)\right)$$

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http://blog.innotas.com/wp-content/uploads/2015/08/chicken<u>-or-egg-cropped1.jpg</u>



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Expectation Maximization (EM)

0. Assume *some* value for your parameters

Two step, iterative algorithm

1. E-step: count under uncertainty (compute expectations)

2. M-step: maximize log-likelihood, assuming these uncertain counts

Expectation Maximization (EM): E-step

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Expectation Maximization (EM): M-step

0. Assume *some* value for your parameters

Two step, iterative algorithm

1. E-step: count under uncertainty, assuming these parameters

2. M-step: maximize log-likelihood, assuming these uncertain counts $p^{(t+1)}(z)$

$$p^{(t)}(z)$$



the average log-likelihood of our Complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

maximize the average log-likelihood of our complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

$\max_{\theta} \mathbb{E}_{z \sim p_{\theta}(t)}(\cdot|w) [\log p_{\theta}(z,w)]$

maximize the average log-likelihood of our complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

$\max_{\theta} \mathbb{E}_{\boldsymbol{Z}} \sim p_{\theta}(t)(\cdot|w) [\log p_{\theta}(\boldsymbol{Z}, w)]$

maximize the average log-likelihood of our complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

 $\max_{\theta} \mathbb{E}_{Z} \sim p_{\theta}(t)(\cdot|w) \left[\log p_{\theta}(Z, w) \right]$ posterior distribution

maximize the average log-likelihood of our complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

 $\max_{\theta} \mathbb{E}_{Z} \sim p_{\theta}(t)(\cdot|w) \left[\log p_{\theta}(Z, w) \right]$

new parameters

posterior distribution

new parameters

maximize the average log-likelihood of our complete data (z, w), averaged across all z and according to how likely our *current* model thinks z is

 $\max_{\theta} \mathbb{E}_{Z} \sim p_{\theta}(t)(\cdot|w) \left[\log p_{\theta}(Z,w) \right]$ *new* parameters posterior distribution *new* parameters

E-step: count under uncertainty M-step: maximize log-likelihood

NO labeled data:

- human annotated
- relatively small/few examples



EM/generative models in this case can be seen as a type of clustering

- raw; not annotated
- plentiful

??? X ??? ??? ??? ??? ??? Х ??? ???

labeled data:

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Imagine three coins







Flip 1st coin (penny)

If heads: flip 2nd coin (dollar coin)

If tails: flip 3rd coin (dime)

Imagine three coins



Flip 1st coin (penny) <

If heads: flip 2nd coin (dollar coin) only observe these (record heads vs. tails outcome)

Imagine three coins







Flip 1st coin (penny) ← unobserved: part of speech? genre?

If heads: flip 2nd coin (dollar coin)

If tails: flip 3rd coin (dime)

observed: *a*, *b*, *e*, etc. We run the code, vs. The *run* failed

Imagine three coins







Flip 1st coin (penny)

$$p(heads) = \lambda$$

 $p(\text{tails}) = 1 - \lambda$

If heads: flip 2^{nd} coin (dollar coin) $p(heads) = \gamma$ If tails: flip 3^{rd} coin (dime)

- $p(tails) = 1 \gamma$
- $p(\text{heads}) = \psi$ $p(\text{tails}) = 1 \psi$

Imagine three coins







 $p(\text{heads}) = \lambda$ $p(\text{heads}) = \gamma$ $p(\text{heads}) = \psi$ $p(\text{tails}) = 1 - \lambda$ $p(\text{tails}) = 1 - \gamma$ $p(\text{tails}) = 1 - \psi$

Three parameters to estimate: λ , γ , and ψ
$$\begin{array}{l} \textbf{Generative Story for Three Coins} \\ p(w_1, w_2, \dots, w_N) = p(w_1)p(w_2) \cdots p(w_N) = \prod_i p(w_i) \\ & \bigcup_{i \in I} p(w_i) \\ & \bigcup_{i \in I} p(w_i) \\ p(z_1, w_1, z_2, w_2, \dots, z_N, w_N) = p(z_1)p(w_1|z_1) \cdots p(z_N)p(w_N|z_N) \\ & = \prod_i p(w_i|z_i) p(z_i) \\ \hline & & p(\text{heads}) = \lambda \\ p(\text{tails}) = 1 - \lambda \\ & p(\text{heads}) = \gamma \\ p(\text{tails}) = 1 - \gamma \\ \hline & & p(\text{heads}) = 1 - \gamma \\ \hline & & p(\text{heads}) = \psi \\ p(\text{tails}) = 1 - \psi \\ \hline & & p(\text{heads}) = 1 - \psi \\ \hline & p(\text{heads}) = 1 -$$

H H T H T H H T H T T T

If all flips were observed

 $p(heads) = \lambda$ $p(heads) = \gamma$ $p(heads) = \psi$ $p(tails) = 1 - \lambda$ $p(tails) = 1 - \gamma$ $p(tails) = 1 - \psi$

H H T H T H H T H T T T

If all flips were observed

$$p(heads) = \lambda$$
 $p(heads) = \gamma$ $p(heads) = \psi$ $p(tails) = 1 - \lambda$ $p(tails) = 1 - \gamma$ $p(tails) = 1 - \psi$

 $p(\text{heads}) = \frac{4}{6} \qquad p(\text{heads}) = \frac{1}{4} \qquad p(\text{heads}) = \frac{1}{2}$ $p(\text{tails}) = \frac{2}{6} \qquad p(\text{tails}) = \frac{3}{4} \qquad p(\text{tails}) = \frac{1}{2}$

H H T H T H H T H T T T

But not all flips are observed \rightarrow set parameter values

 $p(heads) = \lambda = .6$ p(heads) = .8p(heads) = .6p(tails) = .4p(tails) = .2p(tails) = .4

H H T H T H H T H T T T

But not all flips are observed \rightarrow set parameter values

 $p(\text{heads}) = \lambda = .6$ p(heads) = .8 p(heads) = .6p(tails) = .4 p(tails) = .2 p(tails) = .4

Use these values to compute posteriors $p(\text{heads} \mid \text{observed item H}) = \frac{p(\text{heads \& H})}{p(\text{H})}$ $p(\text{heads} \mid \text{observed item T}) = \frac{p(\text{heads \& T})}{p(\text{T})}$

HHTHTH HTHTT

But not all flips are observed \rightarrow set parameter values

 $p(\text{heads}) = \lambda = .6$ p(heads) = .8 p(heads) = .6p(tails) = .4 p(tails) = .2 p(tails) = .4

Use these values to compute posteriors

 $p(\text{heads} \mid \text{observed item H}) = \frac{p(H \mid \text{heads})p(\text{heads})}{p(H)}$ marginal likelihood

H H T H T H H T H T T T

But not all flips are observed \rightarrow set parameter values

 $p(heads) = \lambda = .6$ p(heads) = .8p(heads) = .6p(tails) = .4p(tails) = .2p(tails) = .4

Use these values to compute posteriors

$$p(\text{heads} \mid \text{observed item H}) = \frac{p(H \mid \text{heads})p(\text{heads})}{p(H)}$$

p(H | heads) = .8 p(T | heads) = .2

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But not all flips are observed \rightarrow set parameter values

 $p(heads) = \lambda = .6$ p(heads) = .8p(heads) = .6p(tails) = .4p(tails) = .2p(tails) = .4

Use these values to compute posteriors

$$p(\text{heads} \mid \text{observed item H}) = \frac{p(H \mid \text{heads})p(\text{heads})}{p(H)}$$

p(H | heads) = .8 p(T | heads) = .2

p(H) = p(H | heads) * p(heads) + p(H | tails) * p(tails)= .8 * .6 + .6 * .4

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Use posteriors to update parameters

 $p(\text{heads} \mid \text{obs. H}) = \frac{p(\text{H} \mid \text{heads})p(\text{heads})}{p(\text{H})}$ $= \frac{.8 * .6}{.8 * .6 + .6 * .4} \approx 0.667$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.2 * .6}{.2 * .6 + .6 * .4} \approx 0.334$

Q: Is p(heads | obs. H) + p(heads | obs. T) = 1?

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Use posteriors to update parameters

 $p(\text{heads} \mid \text{obs. H}) = \frac{p(\text{H} \mid \text{heads})p(\text{heads})}{p(\text{H})}$ $= \frac{.8 * .6}{.8 * .6 + .6 * .4} \approx 0.667$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.2 * .6}{.2 * .6 + .6 * .4} \approx 0.334$

Q: Is p(heads | obs. H) + p(heads | obs. T) = 1?

A: No.

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Use posteriors to update parameters

 $p(\text{heads} \mid \text{obs. H}) = \frac{p(\text{H} \mid \text{heads})p(\text{heads})}{p(\text{H})}$ $= \frac{.8 * .6}{.8 * .6 + .6 * .4} \approx 0.667$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.2 * .6}{.2 * .6 + .6 * .4} \approx 0.334$

(in general, p(heads | obs. H) and p(heads | obs. T) do NOT sum to 1)

fully observed setting $p(\text{heads}) = \frac{\# \text{ heads from penny}}{\# \text{ total flips of penny}}$

our setting: partially-observed

$$p(heads) = \frac{\# expected heads}{\# total flips of penny}$$

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Use posteriors to update parameters

 $p(\text{heads} \mid \text{obs. H}) = \frac{p(\text{H} \mid \text{heads})p(\text{heads})}{p(\text{H})}$ $= \frac{.8 * .6}{.8 * .6 + .6 * .4} \approx 0.667$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.2 * .6}{.2 * .6 + .6 * .4} \approx 0.334$

$$p^{(t+1)}(\text{heads}) = \frac{\# \text{ expected heads from penny}}{\# \text{ total flips of penny}}$$
$$= \frac{\mathbb{E}_{p^{(t)}}[\# \text{ expected heads from penny}]}{\# \text{ total flips of penny}}$$

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Use posteriors to update parameters

 $p(\text{heads} \mid \text{obs. H}) = \frac{p(\text{H} \mid \text{heads})p(\text{heads})}{p(\text{H})}$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.8 * .6}{.8 * .6 + .6 * .4} \approx 0.667$ $p(\text{heads} \mid \text{obs. T}) = \frac{p(\text{T} \mid \text{heads})p(\text{heads})}{p(\text{T})}$ $= \frac{.2 * .6}{.2 * .6 + .6 * .4} \approx 0.334$

our setting: partiallyobserved

$$p^{(t+1)}(\text{heads}) = \frac{\# \text{ expected heads from penny}}{\# \text{ total flips of penny}}$$
$$= \frac{\mathbb{E}_{p^{(t)}}[\# \text{ expected heads from penny}]}{\# \text{ total flips of penny}}$$
$$= \frac{2 * p(\text{heads} \mid \text{obs. H}) + 4 * p(\text{heads} \mid \text{obs. T})}{6}$$
$$\approx 0.444$$

Expectation Maximization (EM)

0. Assume *some* value for your parameters

Two step, iterative algorithm:

1. E-step: count under uncertainty (compute expectations)

2. M-step: maximize log-likelihood, assuming these uncertain counts

Outline

Latent and probabilistic modeling Generative Modeling Example 1: A Model of Rolling a Die Example 2: A Model of Conditional Die Rolls

EM (Expectation Maximization) Basic idea Three coins example Why EM works

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \max$ is a log-likelihood of observed data X

 $C(\theta) =$ log-likelihood of complete data (X,Y)

 $\mathcal{P}(\theta) = \text{posterior log-likelihood of}$ incomplete data Y

what do $\mathcal{C}, \mathcal{M}, \mathcal{P}$ look like?

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y)

$$\mathcal{C}(\theta) = \sum_{i} \log p(x_i, y_i)$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y)

$$\mathcal{C}(\theta) = \sum_{i} \log p(x_i, y_i)$$

$$\mathcal{M}(\theta) = \sum_{i} \log p(x_i) = \sum_{i} \log \sum_{k} p(x_i, y = k)$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y)

$$\mathcal{C}(\theta) = \sum_{i} \log p(x_i, y_i)$$

$$\mathcal{M}(\theta) = \sum_{i} \log p(x_i) = \sum_{i} \log \sum_{k} p(x_i, y = k)$$

$$\mathcal{P}(\theta) = \sum_{i} \log p(y_i | x_i)$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \max$ marginal log-likelihood of
observed data X

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of incomplete data Y

 $p_{\theta}(X) = \frac{p_{\theta}(X, Y)}{p_{\theta}(Y \mid X)}$



definition of conditional probability



X: observed dataY: unobserved data $\mathcal{M}(\theta) = \max$ in al log-likelihood of
observed data X

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of incomplete data Y

$$p_{\theta}(Y \mid X) = \frac{p_{\theta}(X, Y)}{p_{\theta}(X)} \quad \Longrightarrow \quad p_{\theta}(X) = \frac{p_{\theta}(X, Y)}{p_{\theta}(Y \mid X)}$$

 $\mathcal{M}(\theta) = \mathcal{C}(\theta) - \mathcal{P}(\theta)$

 $\mathbb{E}_{Y \sim \theta^{(t)}}[\mathcal{M}(\theta)|X] = \mathbb{E}_{Y \sim \theta^{(t)}}[\mathcal{C}(\theta)|X] - \mathbb{E}_{Y \sim \theta^{(t)}}[\mathcal{P}(\theta)|X]$

take a conditional expectation (why? we'll cover this more in variational inference)

X: observed dataY: unobserved dataC $\mathcal{M}(\theta) = \max$ in al log-likelihood of
observed data X $\mathcal{M}(\theta) = \max$

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of incomplete data Y

$$p_{\theta}(Y \mid X) = \frac{p_{\theta}(X, Y)}{p_{\theta}(X)} \qquad \Longrightarrow \qquad p_{\theta}(X) = \frac{p_{\theta}(X, Y)}{p_{\theta}(Y \mid X)}$$

 $\mathcal{M}(\theta) = \mathcal{C}(\theta) - \mathcal{P}(\theta)$ $\mathbb{E}_{Y \sim \theta}(t) [\mathcal{M}(\theta) | X] = \mathbb{E}_{Y \sim \theta}(t) [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta}(t) [\mathcal{P}(\theta) | X]$ $\mathcal{M}(\theta) = \mathbb{E}_{Y \sim \theta}(t) [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta}(t) [\mathcal{P}(\theta) | X]$ $\overset{\mathcal{M}(\theta)}{\text{sums over } Y} \qquad \mathcal{M}(\theta) = \sum_{i} \log p(x_i) = \sum_{i} \log \sum_{k} p(x_i, y = k)$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y)

$$\mathcal{M}(\theta) = \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{P}(\theta) | X]$$

$$\mathbb{E}_{Y \sim \theta^{(t)}}[\mathcal{C}(\theta)|X] = \sum_{i} \sum_{k} p_{\theta^{(t)}}(y = k | x_i) \log p(x_i, y = k)$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of

incomplete data Y

$$\mathcal{M}(\theta) = \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{P}(\theta) | X]$$
$$Q(\theta, \theta^{(t)}) \qquad R(\theta, \theta^{(t)})$$

Let θ^* be the value that maximizes $Q(\theta, \theta^{(t)})$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of

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$$\mathcal{M}(\theta) = \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{P}(\theta) | X]$$
$$Q(\theta, \theta^{(t)}) \qquad R(\theta, \theta^{(t)})$$

Let θ^* be the value that maximizes $Q(\theta, \theta^{(t)})$

$$\mathcal{M}(\theta^*) - \mathcal{M}\left(\theta^{(t)}\right) = \left(Q\left(\theta^*, \theta^{(t)}\right) - Q(\theta^{(t)}, \theta^{(t)})\right) - \left(R\left(\theta^*, \theta^{(t)}\right) - R(\theta^{(t)}, \theta^{(t)})\right)$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of

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$$\mathcal{M}(\theta^*) - \mathcal{M}(\theta^{(t)}) = \left(Q(\theta^*, \theta^{(t)}) - Q(\theta^{(t)}, \theta^{(t)})\right) - \left(R(\theta^*, \theta^{(t)}) - R(\theta^{(t)}, \theta^{(t)})\right)$$
$$\geq 0 \qquad \leq 0 \text{ (we'll see why with Jensen's inequality, in variational inference)}$$

X: observed dataY: unobserved data $\mathcal{M}(\theta) = \text{marginal log-likelihood of observed data X}$

 $C(\theta) =$ log-likelihood of complete data (X,Y) $\mathcal{P}(\theta) =$ posterior log-likelihood of

incomplete data Y

$$\mathcal{M}(\theta) = \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{C}(\theta) | X] - \mathbb{E}_{Y \sim \theta^{(t)}} [\mathcal{P}(\theta) | X]$$
$$Q(\theta, \theta^{(t)}) \qquad R(\theta, \theta^{(t)})$$

Let θ^* be the value that maximizes $Q(\theta, \theta^{(t)})$

 $\mathcal{M}(\theta^*) - \mathcal{M}\left(\theta^{(t)}\right) = \left(Q\left(\theta^*, \theta^{(t)}\right) - Q(\theta^{(t)}, \theta^{(t)})\right) - \left(R\left(\theta^*, \theta^{(t)}\right) - R(\theta^{(t)}, \theta^{(t)})\right)$

$$\mathcal{M}(\theta^*) - \mathcal{M}(\theta^{(t)}) \ge 0$$

EM does not decrease the marginal log-likelihood

Generalized EM

Partial M step: find a θ that simply increases, rather than *maximizes*, Q

Partial E step: only consider *some* of the variables (an online learning algorithm)

EM has its pitfalls

Objective is not convex → converge to a bad local optimum

Computing expectations can be hard: the E-step could require clever algorithms

How well does log-likelihood correlate with an end task?

A Maximization-Maximization Procedure



we'll see this again with variational inference



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