

Predicting Real-valued outputs: an introduction to Regression

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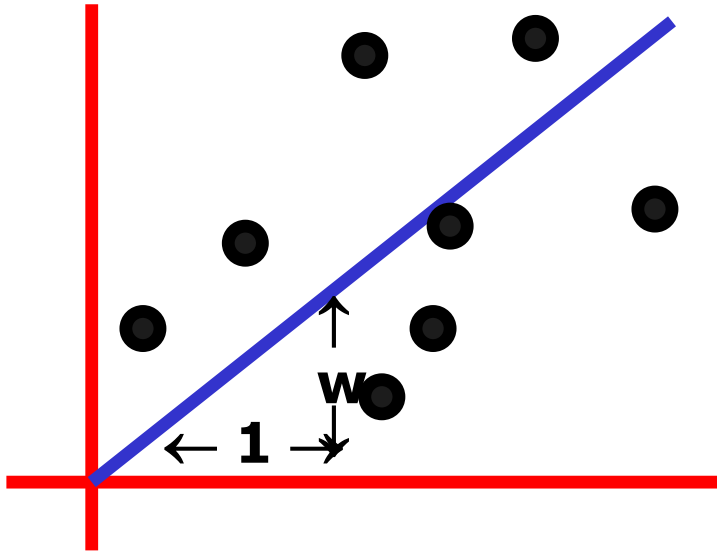
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This is reordered material
from the Neural Nets
lecture and the "Favorite
Regression Algorithms"
lecture

Single- Parameter Linear Regression

Linear Regression

DATASET



inputs	outputs
$x_1 = 1$	$y_1 = 1$
$x_2 = 3$	$y_2 = 2.2$
$x_3 = 2$	$y_3 = 2$
$x_4 = 1.5$	$y_4 = 1.9$
$x_5 = 4$	$y_5 = 3.1$

Linear regression assumes that the expected value of the output given an input, $E[y/x]$, is linear.

Simplest case: $\text{Out}(x) = wx$ for some unknown w .

Given the data, we can estimate w .

1-parameter linear regression

Assume that the data is formed by

$$y_i = wx_i + \text{noise}_i$$

where...

- the noise signals are independent
- the noise has a normal distribution with mean 0 and unknown variance σ^2

$p(y|w,x)$ has a normal distribution with

- mean wx
- variance σ^2

Bayesian Linear Regression

$$p(y|w,x) = \text{Normal}(\text{mean } wx, \text{var } \sigma^2)$$

We have a set of datapoints (x_1, y_1) (x_2, y_2) ... (x_n, y_n) which are **EVIDENCE** about w .

We want to infer w from the data.

$$p(w|x_1, x_2, x_3, \dots, x_n, y_1, y_2, \dots, y_n)$$

- You can use **BAYES** rule to work out a posterior distribution for w given the data.
- Or you could do Maximum Likelihood Estimation

Maximum likelihood estimation of w

Asks the question:

“For which value of w is this data most likely to have happened?”

\Leftrightarrow

For what w is

$p(y_1, y_2, \dots, y_n | x_1, x_2, x_3, \dots, x_n, w)$ maximized?

\Leftrightarrow

For what w is

$\prod_{i=1}^n p(y_i | w, x_i)$ maximized

For what w is

$$\prod_{i=1}^n p(y_i | w, x_i) \text{ maximized}$$

For what w is

$$\prod_{i=1}^n \exp\left(-\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2\right) \text{ maximized?}$$

For what w is

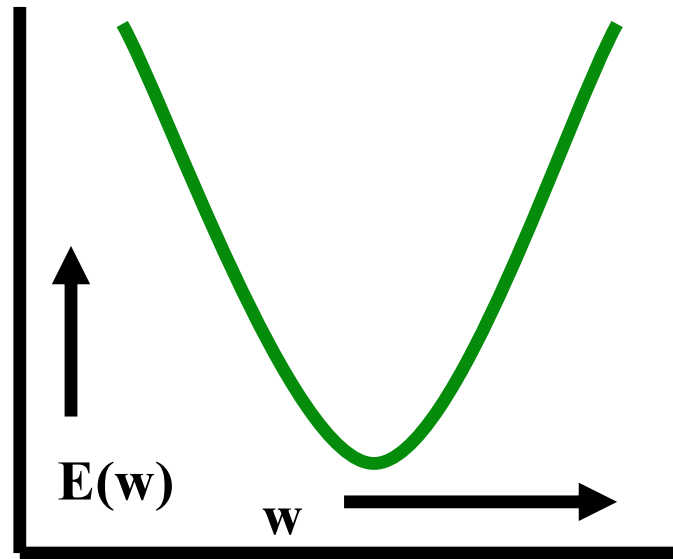
$$\sum_{i=1}^n -\frac{1}{2}\left(\frac{y_i - wx_i}{\sigma}\right)^2 \text{ maximized?}$$

For what w is

$$\sum_{i=1}^n (y_i - wx_i)^2 \text{ minimized?}$$

Linear Regression

The maximum likelihood w is the one that minimizes sum-of-squares of residuals



$$\begin{aligned} E &= \sum_i (y_i - wx_i)^2 \\ &= \sum_i y_i^2 - (2 \sum x_i y_i)w + (\sum x_i^2)w^2 \end{aligned}$$

We want to minimize a quadratic function of w .

Linear Regression

Easy to show the sum of squares is minimized when

$$w = \frac{\sum x_i y_i}{\sum x_i^2}$$

The maximum likelihood model is $\text{Out}(x) = wx$

We can use it for prediction

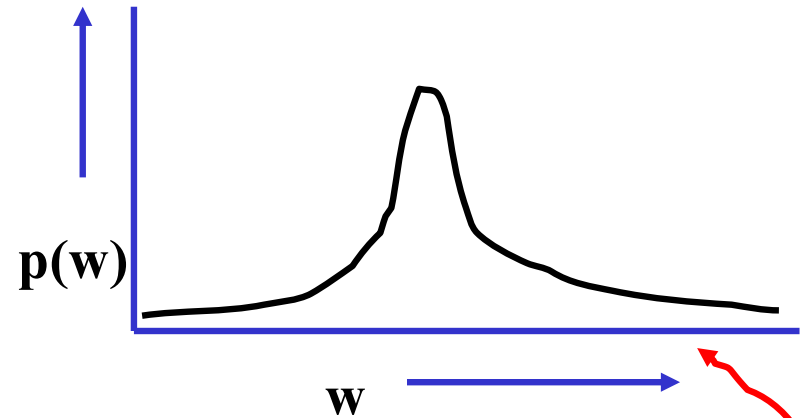
Linear Regression

Easy to show the sum of squares is minimized when

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Note: In Bayesian stats you'd have ended up with a prob dist of w

And predictions would have given a prob dist of expected output

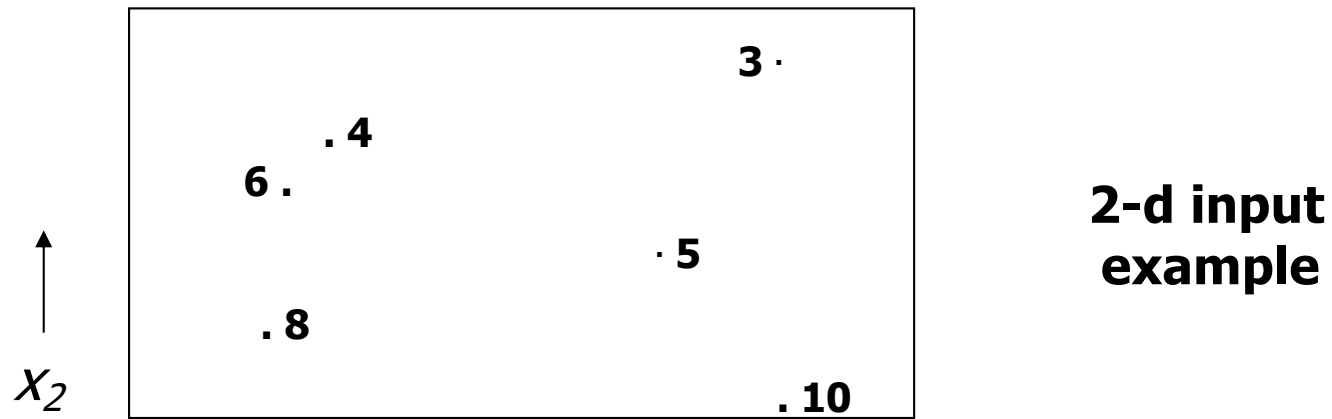
Often useful to know your confidence.

Max likelihood can give some kinds of confidence too.

Multivariate Linear Regression

Multivariate Regression

What if the inputs are vectors?



Dataset has form

$$\begin{array}{l} \mathbf{x}_1 \\ \mathbf{x}_2 \\ \mathbf{x}_3 \\ \vdots \\ \mathbf{x}_R \end{array} \quad \begin{array}{l} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_R \end{array}$$

Multivariate Regression

Write matrix X and Y thus:

$$\mathbf{X} = \begin{bmatrix} \dots \mathbf{x}_1 \dots \\ \dots \mathbf{x}_2 \dots \\ \dots \mathbf{x}_M \dots \\ \dots \mathbf{x}_R \dots \end{bmatrix} = \begin{bmatrix} x_{11} & x_{12} & \dots & x_{1m} \\ x_{21} & x_{22} & \dots & x_{2m} \\ \dots & \dots & \dots & \dots \\ x_{R1} & x_{R2} & \dots & x_{Rm} \end{bmatrix} \quad \mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \dots \\ y_R \end{bmatrix}$$

(there are R datapoints. Each input has m components)

The linear regression model assumes a vector \mathbf{w} such that

$$\text{Out}(\mathbf{x}) = \mathbf{w}^T \mathbf{x} = w_1 x[1] + w_2 x[2] + \dots + w_m x[D]$$

The max. likelihood \mathbf{w} is $\mathbf{w} = (\mathbf{X}^T \mathbf{X})^{-1} (\mathbf{X}^T \mathbf{Y})$

Multivariate Regression

Write matrix X and Y thus:

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(there are R datapoints. Each input

**IMPORTANT EXERCISE:
PROVE IT !!!!!**

The linear regression model assumes a vector \mathbf{w} such that

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The max. likelihood \mathbf{w} is $\mathbf{w} = (\mathbf{X}^T \mathbf{X})^{-1} (\mathbf{X}^T \mathbf{Y})$

Multivariate Regression (con't)

The max. likelihood \mathbf{w} is $\mathbf{w} = (X^T X)^{-1} (X^T Y)$

$X^T X$ is an $m \times m$ matrix: i, j 'th elt is $\sum_{k=1}^R x_{ki} x_{kj}$

$X^T Y$ is an m -element vector: i 'th elt $\sum_{k=1}^R x_{ki} y_k$

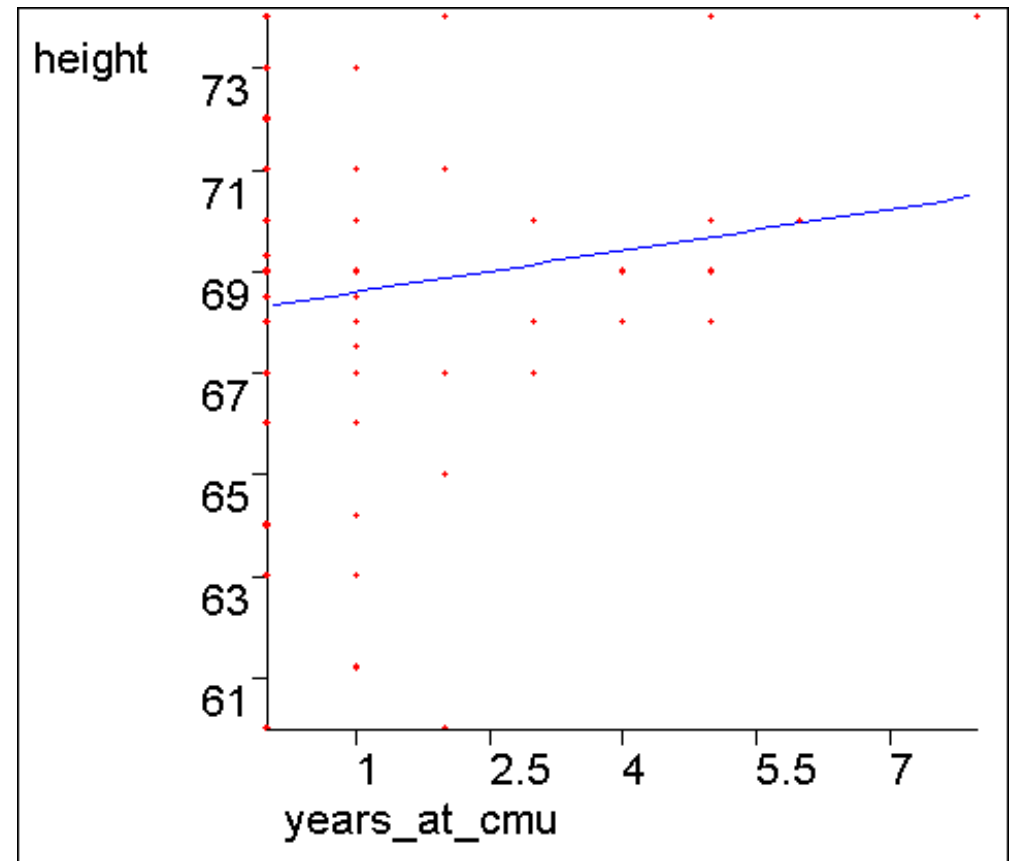
Constant Term in Linear Regression

What about a constant term?

We may expect linear data that does not go through the origin.

Statisticians and Neural Net Folks all agree on a simple obvious hack.

Can you guess??



The constant term

- The trick is to create a fake input “ X_0 ” that always takes the value 1

X_1	X_2	Y
2	4	16
3	4	17
5	5	20

Before:

$$Y = w_1 X_1 + w_2 X_2$$

...has to be a poor model

X_0	X_1	X_2	Y
1	2	4	16
1	3	4	17
1	5	5	20

After:

$$Y = w_0 X_0 + w_1 X_1 + w_2 X_2$$

$$= w_0 + w_1 X_1 + w_2 X_2$$

...has a fine constant term

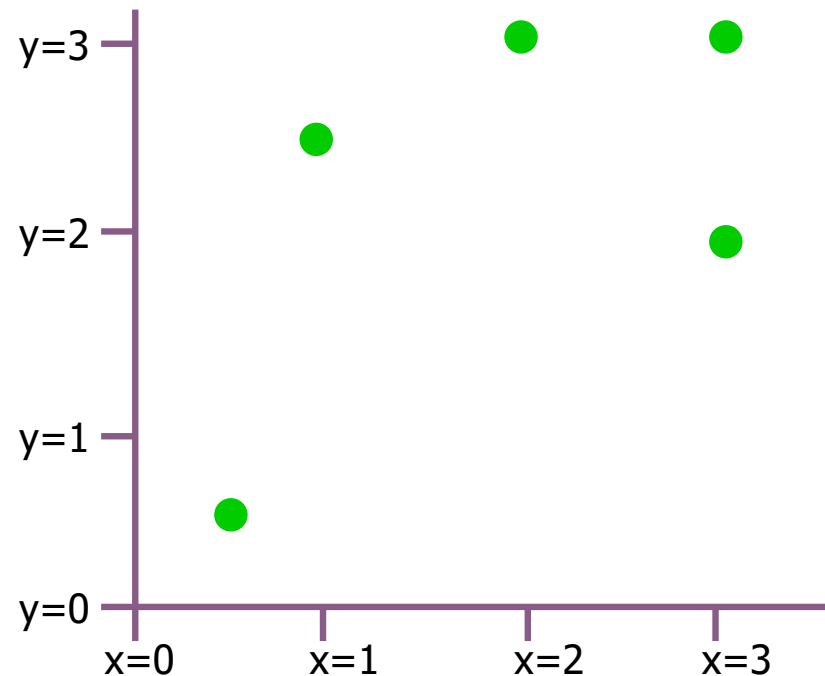
In this example, You should be able to see the MLE w_0 , w_1 and w_2 by inspection

Non-linear Regression

Non-linear Regression

- Suppose you know that y is related to a function of x in such a way that the predicted values have a non-linear dependence on w , e.g:

X_i	Y_i
$1/2$	$1/2$
1	2.5
2	3
3	2
3	3



Assume $y_i \sim N(\sqrt{w + x_i}, \sigma^2)$

What's the MLE estimate of w ?

Non-linear MLE estimation

$$\operatorname{argmax}_w \log p(y_1, y_2, \dots, y_R \mid x_1, x_2, \dots, x_R, \sigma, w) =$$

$$\operatorname{argmin}_w \sum_{i=1}^R (y_i - \sqrt{w + x_i})^2 =$$

Assuming i.i.d. and then plugging in equation for Gaussian and simplifying.

$$\left(w \text{ such that } \sum_{i=1}^R \frac{y_i - \sqrt{w + x_i}}{\sqrt{w + x_i}} = 0 \right) =$$

Setting dLL/dw equal to zero

Non-linear MLE estimation

$$\operatorname{argmax}_w \log p(y_1, y_2, \dots, y_R \mid x_1, x_2, \dots, x_R, \sigma, w) =$$

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Setting dLL/dw equal to zero



We're down the algebraic toilet

So guess what we do?

Non-linear MLE estimation

$$\operatorname{argmax}_w \log p(y_1, y_2, \dots, y_R \mid x_1, x_2, \dots, x_R, \sigma, w) =$$

Common (but not only) approach:
Numerical Solutions:

- Line Search
- Simulated Annealing
- Gradient Descent
- Conjugate Gradient
- Levenberg Marquart
- Newton's Method

Also, special purpose statistical-optimization-specific tricks such as E.M. (See Gaussian Mixtures lecture for introduction)

$$\left. \begin{aligned} \frac{\partial}{\partial w} \log p(y_i \mid x_i, \sigma, w) \\ \frac{\partial}{\partial w} \log p(y_i \mid x_i, \sigma, w) \end{aligned} \right\} =$$

Assuming i.i.d. and then plugging in equation for Gaussian and simplifying.

$$\left. \begin{aligned} \frac{\partial}{\partial w} \log p(y_i \mid x_i, \sigma, w) \\ \frac{\partial}{\partial w} \log p(y_i \mid x_i, \sigma, w) \\ \frac{\partial}{\partial w} \log p(y_i \mid x_i, \sigma, w) \end{aligned} \right) = 0 =$$

Setting dLL/dw equal to zero

We're down the algebraic toilet



So guess what we do?

Polynomial Regression

Polynomial Regression

So far we've mainly been dealing with linear regression

X_1	X_2	Y
3	2	7
1	1	3
\vdots	\vdots	\vdots

$\mathbf{X} = \begin{bmatrix} 3 & 2 \\ 1 & 1 \\ \vdots & \vdots \end{bmatrix}$ $\mathbf{y} = \begin{bmatrix} 7 \\ 3 \\ \vdots \end{bmatrix}$

$(3 \ 2) \dots$ $y_1 = 7 \dots$

$\mathbf{Z} = \begin{bmatrix} 1 & 3 & 2 \\ 1 & 1 & 1 \\ \vdots & \vdots & \vdots \end{bmatrix}$ $\mathbf{y} = \begin{bmatrix} 7 \\ 3 \\ \vdots \end{bmatrix}$

$\mathbf{z}_1 = (1, 3, 2) \dots$ $y_1 = 7 \dots$

$\mathbf{z}_k = (1, x_{k1}, x_{k2})$

$$\boldsymbol{\beta} = (\mathbf{Z}^T \mathbf{Z})^{-1} (\mathbf{Z}^T \mathbf{y})$$

$$y^{est} = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

Quadratic Regression

It's trivial to do linear fits of fixed nonlinear basis functions

X_1	X_2	Y
3	2	7
1	1	3
\vdots	\vdots	\vdots

$\mathbf{x} =$	3	2
	1	1
	\vdots	\vdots

$\mathbf{y} =$	7
	3
	\vdots

$y_1 = 7..$

$\mathbf{z} =$	1	3	2	9	6	4
	1	1	1	1	1	1
	\vdots					\vdots

$\mathbf{y} =$	7
	3
	\vdots

$\mathbf{z} = (1, X_1, X_2, X_1^2, X_1X_2, X_2^2)$

$$\beta = (\mathbf{Z}^T \mathbf{Z})^{-1} (\mathbf{Z}^T \mathbf{y})$$

$$y^{est} = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1^2 + \beta_4 X_1 X_2 + \beta_5 X_2^2$$

Quadratic Regression

It's trivial

X_1	X_2
3	2
1	1
\vdots	\vdots

$\mathbf{Z} =$

1
1
\vdots

$\mathbf{z} = (1$

Each component of a \mathbf{z} vector is called a term.

Each column of the \mathbf{Z} matrix is called a term column

How many terms in a quadratic regression with m inputs?

- 1 constant term
- m linear terms
- $(m+1)\text{-choose-}2 = m(m+1)/2$ quadratic terms
- $(m+2)\text{-choose-}2$ terms in total = $O(m^2)$

Note that solving $\beta = (\mathbf{Z}^T \mathbf{Z})^{-1} (\mathbf{Z}^T \mathbf{y})$ is thus $O(m^6)$

Qth-degree polynomial Regression

X_1	X_2	Y
3	2	7
1	1	3
⋮	⋮	⋮

$\mathbf{X} =$	3	2	$\mathbf{y} =$	7
	1	1		3
	⋮	⋮		⋮

$\mathbf{Z} =$	1	3	2	9	6	...	$\mathbf{y} =$	7
	1	1	1	1	1	...		3
	⋮					...		⋮

\mathbf{z} = (all products of powers of inputs in which sum of powers is q or less,)

$$\beta = (\mathbf{Z}^T \mathbf{Z})^{-1} (\mathbf{Z}^T \mathbf{y})$$

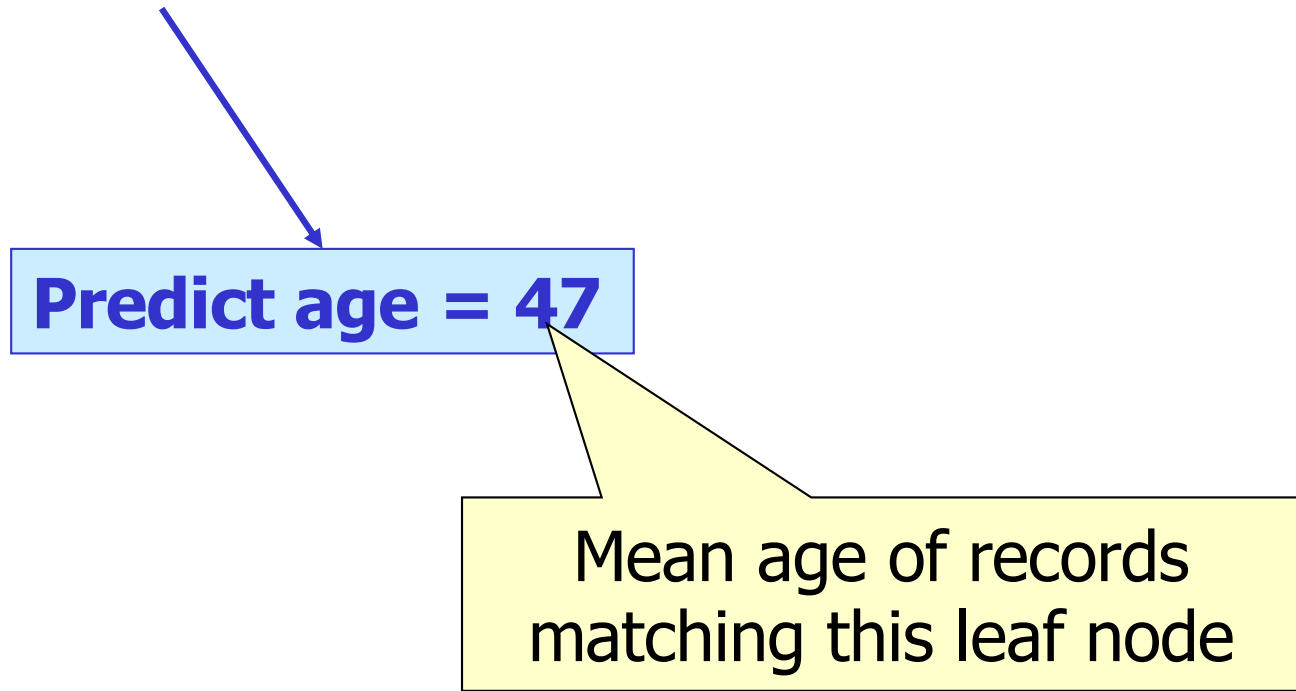
$$y^{est} = \beta_0 + \beta_1 X_1 + \dots$$

Regression Trees

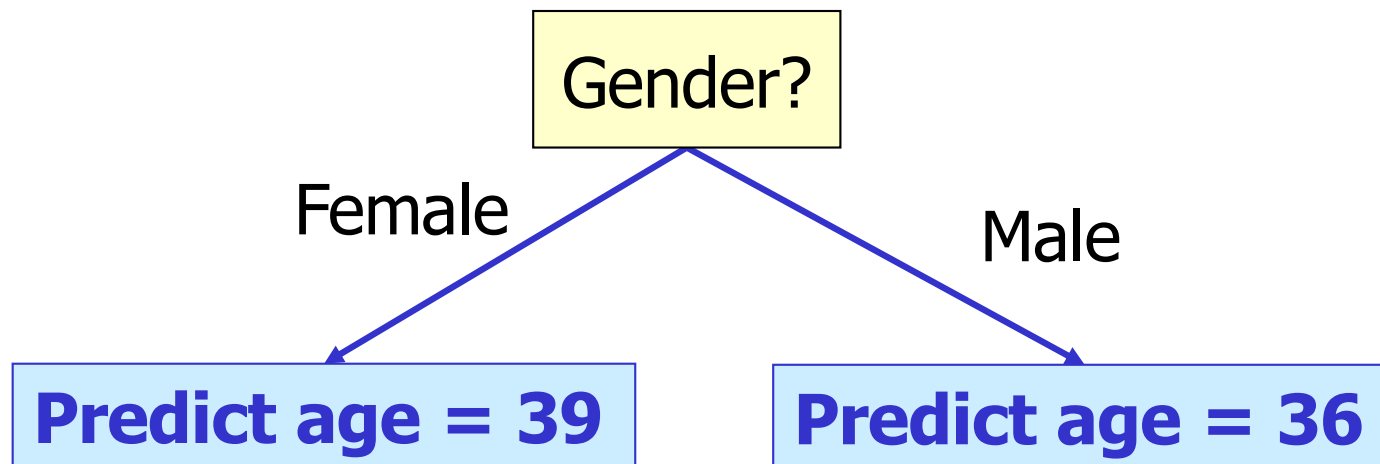
Regression Trees

- “Decision trees for regression”

A regression tree leaf



A one-split regression tree



Choosing the attribute to split on

Gender	Rich?	Num. Children	Num. Beany Babies	Age
Female	No	2	1	38
Male	No	0	0	24
Male	Yes	0	5+	72
:	:	:	:	:

- We can't use information gain.
- What should we use?

Choosing the attribute to split on

Gender	Rich?	Num. Children	Num. Beany Babies	Age
Female	No	2	1	38
Male	No	0	0	24
Male	Yes	0	5+	72
:	:	:	:	:

$MSE(Y|X)$ = The expected squared error if we must predict a record's Y value given only knowledge of the record's X value

If we're told $x=j$, the smallest expected error comes from predicting the mean of the Y-values among those records in which $x=j$. Call this mean quantity $\mu_y^{x=j}$

Then...

$$MSE(Y|X) = \frac{1}{R} \sum_{j=1}^{N_X} \sum_{(k \text{ such that } x_k=j)} (y_k - \mu_y^{x=j})^2$$

Choosing the attribute to split on

Gender	Rich?	Num. Children	Num. Beany Babies	Age
Female	N			
Male	N			
Male	Y			
⋮	⋮			

Regression tree attribute selection: greedily choose the attribute that minimizes $MSE(Y|X)$

Guess what we do about real-valued inputs?

Guess how we prevent overfitting

$MSE(Y|X)$ = The expected squared error if we must predict a record's Y value given only knowledge of the record's X value

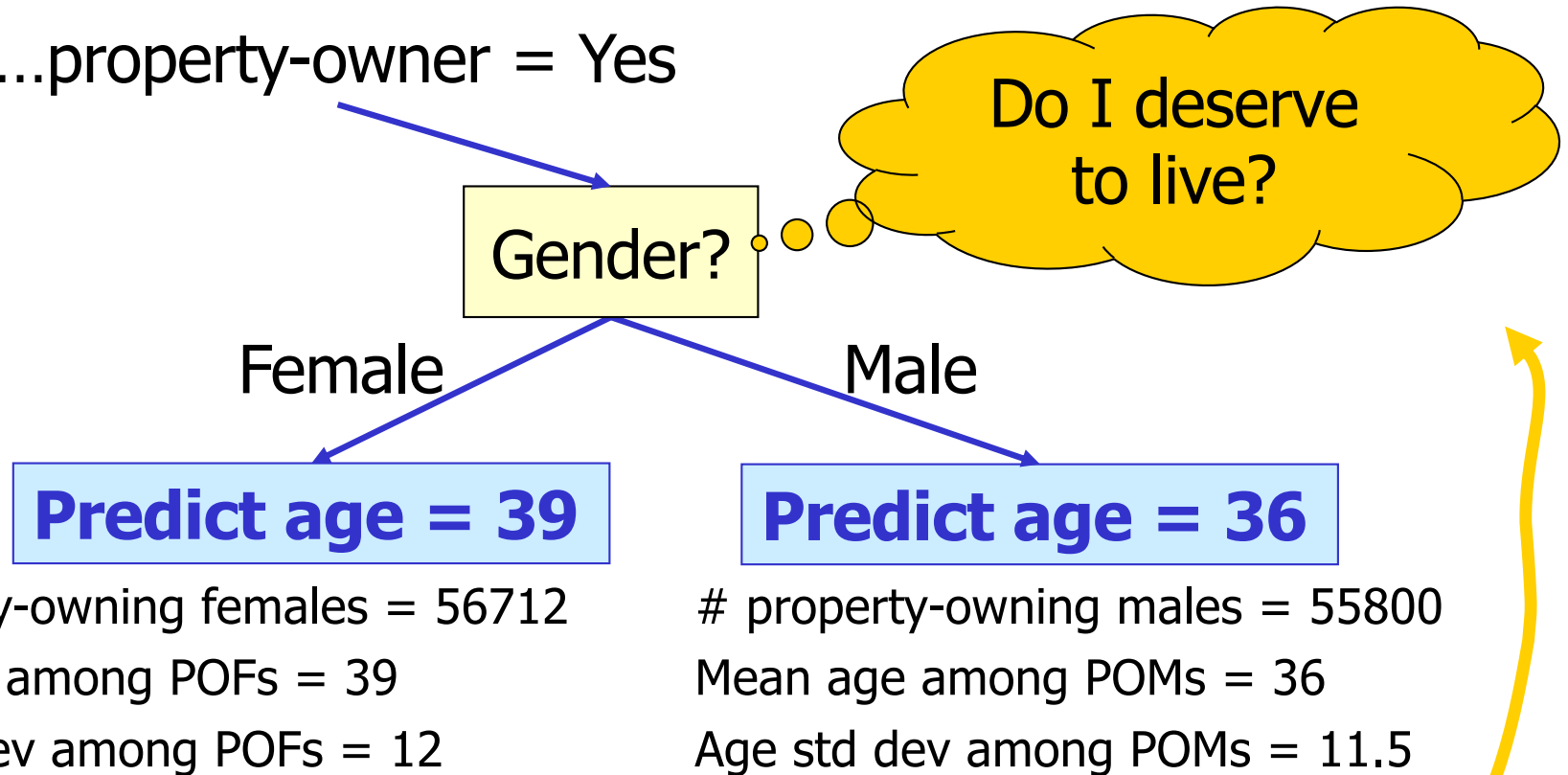
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Then...

$$MSE(Y|X) = \frac{1}{R} \sum_{j=1}^{N_X} \sum_{(k \text{ such that } x_k=j)} (y_k - \mu_y^{x=j})^2$$

Pruning Decision

...property-owner = Yes

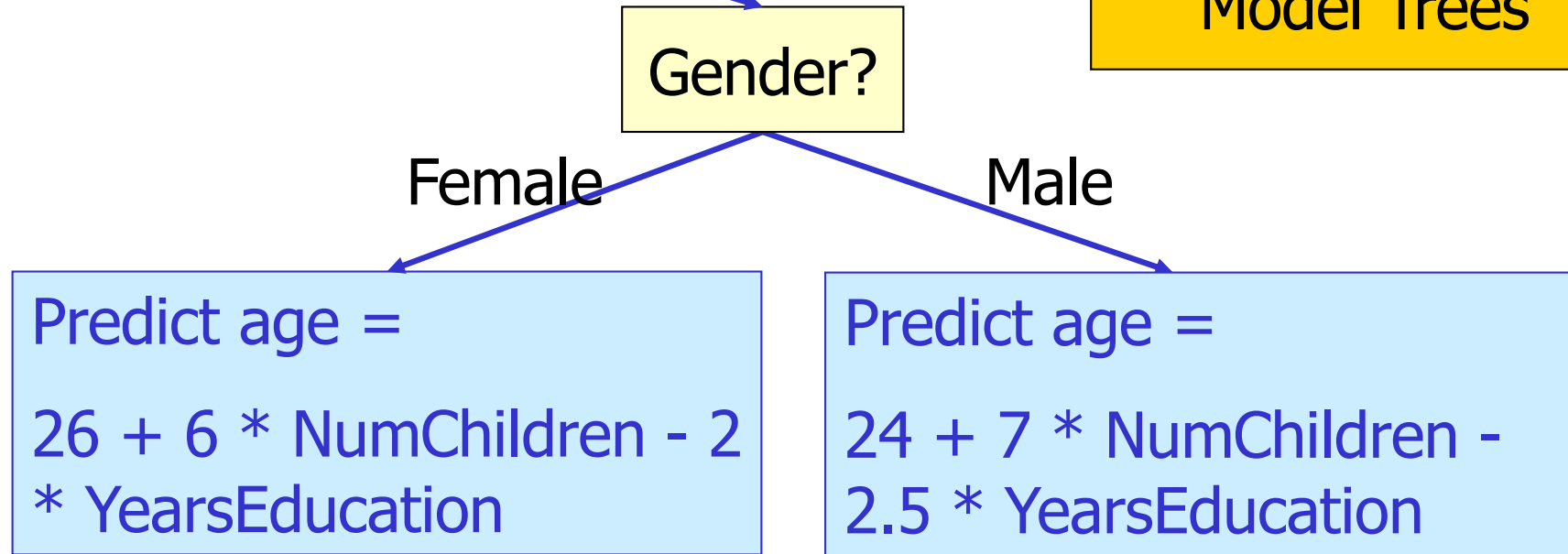


Use a standard Chi-squared test of the null-hypothesis "these two populations have the same mean" and Bob's your uncle.

Linear Regression Trees

...property-owner = Yes

Also known as
"Model Trees"



Leaves contain linear functions (trained using linear regression on all records matching that leaf)

Split attribute chosen to minimize MSE of regressed children.

Pruning with a different Chi-squared

Linear Regression Trees

...property-owner = Yes

Gender?

Female

Predict age =

$26 + 6 * N$

$* \text{YearsE}$

Also known as
"M... Trees"

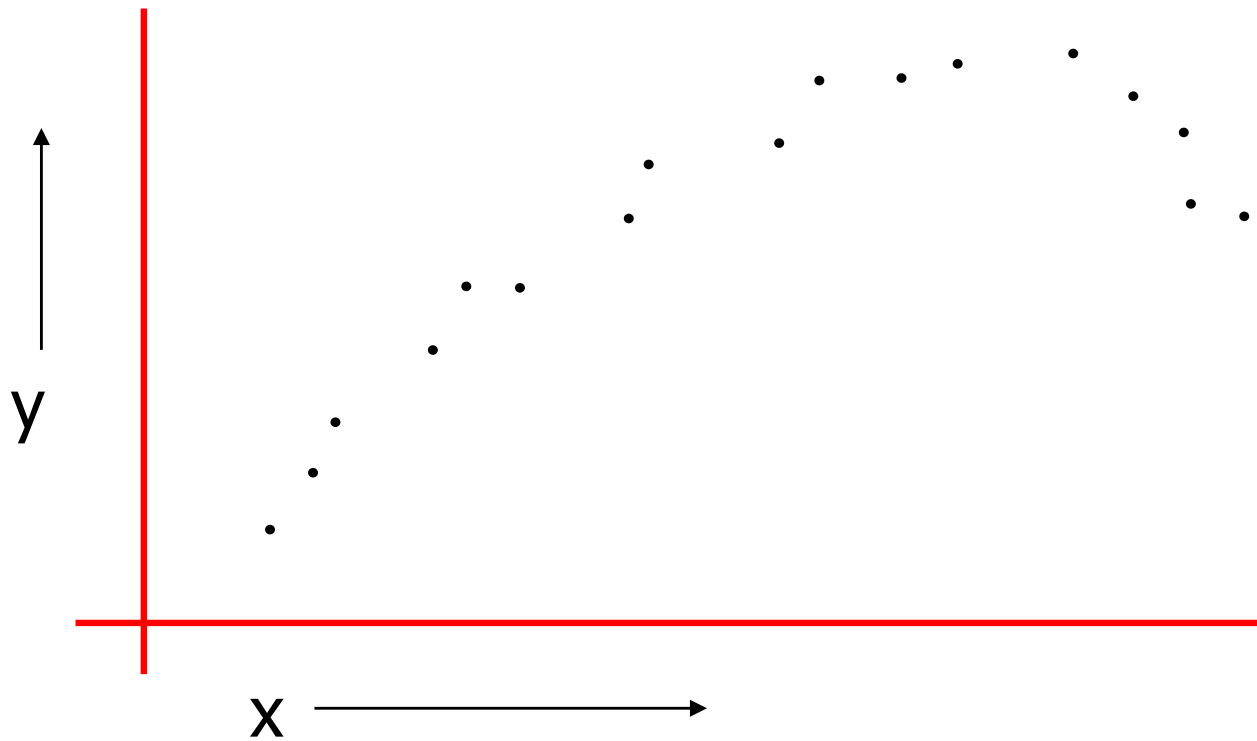
Detail: You typically ignore any categorical attribute that has been tested on higher up in the tree during the regression. But use all untested attributes, and use real-valued attributes even if they've been tested above

Leaves contain linear regression functions (trained on records matching the node) chosen to minimize regression error.

Pruning with a different Chi-squared

Test your understanding

Assuming **regular** regression trees, can you sketch a graph of the fitted function $y^{est}(x)$ over this diagram?



Test your understanding

Assuming **linear** regression trees, can you sketch a graph of the fitted function $y^{est}(x)$ over this diagram?

